# Finite sample valid probabilistic inference on quantile regression 

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## 1- Introduction

Data $Z^{n}=\left\{Z_{i}=\left(X_{i}, Y_{i}\right): i=1, \ldots, n\right\}$ of $n$ covariates-response pairs are iid with distribution P . Nothing is assumed about P .

- $Q_{x}(\tau)=x^{\top} \theta$ is the $\tau$-th quantile of $Y$ given $X=x$.
- Goal: make inferences on $\theta$ that are distribution-free and valid.
- Common solution: Inferences through confidence regions
- Notion of validity is familiar $\rightarrow$ coverage guarantees:

$$
\begin{equation*}
\sup _{P} P^{n}\left\{C_{\alpha}\left(Z^{n}\right) \not \supset \theta(P)\right\} \leq \alpha, \quad \alpha \in[0,1] . \tag{1}
\end{equation*}
$$

## 2 - Probabilistic Inference

Beyond confidence regions, where we can assign degrees of belief $\Pi$ to relevant assertions about $\theta$, e.g., $\theta \in A, A \subseteq \Theta$.

- Validity: control the assignment of high degrees of belief to false assertions:

$$
\begin{equation*}
\sup _{P: \theta(P) \notin A} P^{n}\left\{\Pi_{Z^{n}}(A)>1-\alpha\right\} \leq \alpha, \quad \alpha \in[0,1] . \tag{2}
\end{equation*}
$$

- Bayesian approach? False Confidence Theorem says we need imprecision!


## 3 - Inferential Models

Consider the parametric case where $Z^{n}=\left(Z_{1}, \ldots Z_{n}\right)$ are iid with distribution $P_{\omega}$. IM approach [2] offers valid probabilistic inference for $\omega$.

## Two-step IM construction

1. Choose an appropriate $h:\left(\mathbb{Z}^{n} \times \Omega\right) \rightarrow \mathbb{R}$ that determines a partial ordering of candidate values for $\omega$ given $z^{n}$, e.g.,likelihood ratio:

$$
h\left(z^{n}, \omega\right)=L_{z^{n}}(\omega) / L_{z^{n}}\left(\hat{\omega}_{z^{n}}\right)
$$

2. Compute the possibility contour

$$
\pi_{z^{n}}(\omega)=P_{\omega}^{n}\left\{h\left(Z^{n}, \omega\right) \leq h\left(z^{n}, \omega\right)\right\}
$$

$\pi \rightarrow$ valid probabilistic inference and confidence regions for $\omega$

## 4 - Nonparametric IM for quantile regression

The idea here is to mimic the construction above:

1. Choose an $h$ that orders candidate values for $\theta$ given $z^{n}$
2. Compute the contour

$$
\begin{equation*}
\pi_{z^{n}}(\theta)=P^{n}\left\{h\left(Z^{n}, \theta\right) \leq h\left(z^{n}, \theta\right)\right\}, \quad \theta \in \Theta \tag{3}
\end{equation*}
$$

## Theorem:

- The degrees of belief obtained from (3) are valid in the sense of (2).
- $\left\{\theta \in \Theta: \pi_{z^{n}}(\theta)>\alpha\right\}$ is a valid confidence region in the sense of (1).


## Challenges:

1. What h ? No model $\rightarrow$ no likelihood ratio.
2. How to compute (3)? Recall that P is unknown.

## A possible solution

- In [1], a bootstrap-based IM construction was proposed
- validity is just achieved asymptotically

But we want more! $\rightarrow$ IM that achieves validity for any sample size
Strategy: Choose an $h$ whose distribution is known and independent of unknown quantities, so (3) can be computed! More specifically:

- Find $\gamma\left(\theta, z^{n}\right)$ that is a pivot
- $\mathrm{h} \rightarrow \gamma^{\prime}$ 's probability mass


## 4.1 - An intuitive (but bad) solution

- $\gamma=\sum_{i=1}^{n} I_{(0, \infty)}\left(Y_{i}-x_{i}^{\top} \theta\right) \rightarrow \gamma \sim \operatorname{Bin}(n, 1-\tau) \rightarrow h=\binom{n}{\gamma}(1-\tau)^{\gamma} \tau^{n-\gamma}$
- Very inefficient! For example, for $\tau=0.5$, any line that splits the data in half, e.g., the black, red and blue below, is equally maximally plausible.



## 4.2-A better solution

Let X be discrete with $k$ levels. The idea is to consider the binomials for each level of X separately, and have $h$ as the product of their probability masses.

$$
\begin{equation*}
h=\prod_{i=1}^{k}\binom{n_{i}}{\gamma_{i}}(1-\tau)^{\gamma_{i}} \tau^{n_{i}-\gamma_{i}}, \text { where } \quad \gamma_{i}=\sum_{j=1}^{n_{i}} I_{(0, \infty)}\left(Y_{j}-x_{i} \theta\right) \tag{4}
\end{equation*}
$$

Example: $\mathrm{k}=3, n_{1}=n_{2}=n_{3}=10, \tau=0.5$ :

$$
\text { Figure 1: Data set on the left. } 95 \% \text { confidence region for } \theta \text { on the right. }
$$

If X is continuous, there is no replication of $Y$ for any given $X=x$. But we do have replications of $Y$ in neighborhoods of $X$, so (4) can still be used!

- Form $k$ neighborhoods of $X$
- Consider each one of the $k$ independent binomials separately
- Use the product of their probability masses as the plausibility order $h$

Example: Simulation study with $n=30, \tau=0.3$ and $k=2$ to compare the coverage probabilities and mean length of $95 \%$ interval estimates for the quantile regression coefficients based on the IM and two other methods:

| $\theta$ | IM | Rank | Bayes |
| :---: | :---: | :---: | :---: |
| $\theta_{0}$ | $0.99(1.11)$ | $0.88(0.43)$ | $0.96(0.44)$ |
| $\theta_{1}$ | $0.98(0.48)$ | $0.83(0.19)$ | $0.88(0.18)$ |

## 5 - Open questions

Other pivot options? How to best select the neighborhoods of a continuous $X$ ? Specifically, does the number of neighborhoods and/or the number of replications per neighborhood impact the efficiency of the IM?

## References

[1] L. Cella and R. Martin. Direct and approximately valid probabilistic inference on a class of statistical functionals. International Journal of Approximate Reasoning, 151:205-224, 2022.
[2] R. Martin and C. Liu. Inferential Models: Reasoning with Uncertainty. Monographs in Statistics and Applied Probability Series. Chapman \& Hall/CRC Press, 2015.

